# Exploring the variation in associations between socioeconomic indicators and non-communicable diseases in the Tromsø Study: an algorithmic approach 

SIGBJØRN SVALESTUEN ${ }^{1},{ }^{2}$ (D) EMRE SARI ${ }^{2}$ (D) PETJA LYN LANGHOLZ ${ }^{3}$ (D) \& CHI QUYNH VO ${ }^{4}$ id<br>${ }^{1}$ Department of Social Sciences, UiT The Arctic University of Norway, Tromsø, Norway, ${ }^{2}$ Health Services and Health Economics, NORCE Norwegian Research Centre AS, Tromsø, Norway, ${ }^{3}$ Department of Archaeology, History, Religiuos Studies and Theology, UiT The Arctic University of Norway, Tromsø, Norway and ${ }^{4}$ Department of Community Medicine, UiT The Arctic University of Norway, Tromsø, Norway


#### Abstract

Aims: We contribute to the methodological literature on the assessment of health inequalities by applying an algorithmic approach to evaluate the capabilities of socioeconomic variables in predicting the prevalence of non-communicable diseases in a Norwegian health survey. Methods: We use data from the seventh survey of the population based Tromsø Study (20152016), including 11,074 women and 10,009 men aged 40 years and above. We apply the random forest algorithm to predict four non-communicable disease outcomes (heart attack, cancer, diabetes and stroke) based on information on a number of social root causes and health behaviours. We evaluate our results using the classification error, the mean decrease in accuracy, partial dependence statistics. Results: Results suggest that education, household income and occupation to a variable extent contribute to predicting non-communicable disease outcomes. Prediction misclassification ranges between $25.1 \%$ and $35.4 \%$ depending on the non-communicable diseases under study. Partial dependences reveal mostly expected health gradients, with some examples of complex functional relationships. Out-of-sample model validation shows that predictions translate to new data input. Conclusions: Algorithmic modelling can provide additional empirical detail and metrics for evaluating heterogeneous inequalities in morbidity. The extent to which education, income and occupation contribute to predicting binary non-communicable disease outcomes depends on both non-communicable diseases and socioeconomic indicator. Partial dependences reveal that social gradients in non-communicable disease outcomes vary in shape between combinations of non-communicable disease outcome and socioeconomic status indicator. Misclassification rates highlight the extent of variation within socioeconomic groups, suggesting that future studies may improve predictive accuracy by exploring further subpopulation heterogeneity.


Keywords: Non-communicable disease, socioeconomic status, machine learning, random forest, feature importance, prediction, partial dependence

## Introduction

Decades of studies rooted in traditional data modelling have produced convincing evidence for the existence and persistence of health inequalities. Particular attention has been given to the disease distribution between social groups and the social origins of disease [1-5]. However, issues of model dependence
and the shortcomings of statistical significance as a single evaluation criterion $[6,7]$ suggest that existing metrics of scientific evaluation should be supplemented. Algorithmic modelling can provide such additional metrics by providing indicators of the predictive accuracy of the model in predicting new data input, and reduce model dependence through nonparametric estimation [8].

[^0]Date received 3 March 2023; reviewed 26 March 2024; accepted 9 April 2024

We therefore adopt an algorithmic modelling strategy to explore the variation in association between a set of socioeconomic status (SES) indicators and a set of non-communicable diseases (NCDs). We contribute to discussions on the statistical assessment of socioeconomic correlates of disease by applying an out-of-sample predictive approach in the context of a comprehensive Norwegian health survey. We apply the random forest [9] algorithm on classification problems, and compute variable importance statistics to assess to what extent these factors contribute to correctly predicting a history of NCD outcomes out of sample. We further compute partial dependences between SES indicators and NCD outcomes to assess how education, occupation and income aids the algorithm in making predictions. We discuss the potential benefits of integrating algorithmic modelling into the methodological toolbox of health inequality researchers.

## Methods and materials

## The Tromsø Study

The Tromsø Study is conducted in Tromsø municipality in northern Norway and aims to include large, representative samples of the local population. The invitation of whole birth cohorts and random sampling ensures a balanced representation of both genders, and a demographic closely mirroring the Tromsø population. The survey includes a wide range of variables, including both questionnaires, biological samples and clinical examinations. All inhabitants aged 40 years and older $(N=32,591)$ were invited to Tromsø7 (2015-2016, 65\% participated) [10].

Table I presents summary statistics on all predictors and outcomes applied in the algorithm. The overall sample size was $N=21,083$. We applied simple mean and median imputation to missing values. Our outcomes are four NCDs: diabetes, heart attack, stroke and cancer. These NCDs are leading causes of premature mortality and represent a great burden of disease globally [11]. The outcome measures comprise self-reports of both previous and current disease status. SES was measured using traditional indicators such as level of education, household income and occupational group. Following Olsen et al. [12] we categorised occupations as follows: unskilled (including semi-skilled manual jobs); intermediary (including office, sales, service and care jobs); lower professions requiring tertiary education of up to 3 years; and higher professions including administrative leaders, politicians, or professions requiring at least 4 years of tertiary education.

Education was split into four categories: primary, vocational, tertiary education of less than 4 years and tertiary education of 4 years or more. Household income was measured as eight absolute income categories, ranging from less than NOK150,000 to greater than NOK1,000,000.

Additional predictors were selected to include a broad range of social and demographic correlates of the socioeconomic indicators and health. We included key modifiable risk factors for NCDs such as alcohol consumption, physical activity level, body-mass index and smoking habits [13]. Indicators such as financial conditions in childhood and parental education level aim to capture the impact of early-life conditions on adult health outcomes [4]. We further included indicators of healthcare engagement (frequency of general practitioner and specialist consultations) as research has shown socioeconomic differences in general practitioner service provision [14], and inequalities in private medical specialist utilisation and hospital outpatient care [15]. Further, indicators of financial security and perceptions of occupational status aim to capture health effects related to relative social positioning and psychosocial stress [16]. Marital status was included to differentiate between single and dual household incomes. Finally, we include age and gender as demographic indicators.

## Ethical statement

All participants gave informed written consent, and the study was assessed by the Norwegian Center for Research Data (reference 869500). This study was not defined as health research by the Regional Ethics Committee North and was exempted from the requirement of study preapproval.

## Predicting in and out of sample

Stochastic models imply strong assumptions on the data generating process [17]. Approximating a true functional relationship with simple parametric models such as linear regression introduces some error. The difference between the estimated function and the empirical function is the model bias [18]. Predictive exercises for model validation are a common feature of this strategy in the form of goodness of fit tests and in-sample predictive power [6]. They tend to yield coefficient estimates that give the best in-sample predictions [8].

Supervised machine learning methods search for functions that predict an output for the dependent variable given the independent variable input. Their goal is to achieve a balance between reducing the insample and out-of-sample errors, by searching for
Table I. Summary statistics by sex and totals in the Seventh Tromsø Study (Tromsø7) 2015-2016.

| Variables | Women |  | Men |  | Total |  | Variables | Women |  | Men |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $N$ | Mean/\% | $N$ | Mean/\% | $N$ | Mean/\% |  | $N$ | Mean/\% | $N$ | Mean/\% | $N$ | Mean/\% |
| Age | 11,074 | 57.2 | 10,009 | 57.4 | 21,083 | 57.3 | Consultation GP: Yes | 9283 | 84.6 | 7569 | 76 | 16,852 | 80.5 |
| BMI | 11,038 | 26.9 | 9982 | 27.8 | 21,020 | 27.3 | Consultation, specialist: Yes | 2500 | 23.8 | 1682 | 17.4 | 4182 | 20.8 |
| Education |  |  |  |  |  |  | Physical activity, leisure |  |  |  |  |  |  |
| Primarylower secondary | 2617 | 24.1 | 2179 | 22.2 | 4796 | 23.2 | Reading, watching TV etc. | 1466 | 13.8 | 1506 | 15.4 | 2972 | 14.6 |
| Vocational/upper secondary | 2759 | 25.4 | 2997 | 30.5 | 5756 | 27.8 | Walking, cycling etc. at least $4 \mathrm{~h} /$ week | 6897 | 65 | 4918 | 50.4 | 11,815 | 58 |
| College/university, $<4$ years | 1917 | 17.6 | 2091 | 21.3 | 4008 | 19.4 | Recreational sports, snow shoveling etc. at least $4 \mathrm{~h} /$ week | 2000 | 18.8 | 2951 | 30.2 | 4951 | 24.3 |
| College/university $>4$ years | 3581 | 32.9 | 2564 | 26.1 | 6145 | 29.7 | Hard training, several times/week | 250 | 2.4 | 382 | 3.9 | 632 | 3.1 |
| Household income (NOK thousands) |  |  |  |  |  |  | Employment |  |  |  |  |  |  |
| $<150$ | 134 | 1.3 | 76 | 0.8 | 210 | 1.0 | Works full time | 5694 | 52.2 | 6354 | 64.6 | 12,048 | 58.1 |
| 150-250 | 635 | 6.1 | 355 | 3.6 | 990 | 4.9 | Works part time | 1248 | 11.4 | 414 | 4.2 | 1662 | 8 |
| 251-350 | 911 | 8.7 | 528 | 5.4 | 1439 | 7.1 | Unemployed | 53 | 0.5 | 84 | 0.9 | 137 | 0.7 |
| 351-450 | 1120 | 10.8 | 786 | 8.0 | 1906 | 9.4 | Housekeeping | 103 | 0.9 | 29 | 0.3 | 132 | 0.6 |
| 451-550 | 1319 | 12.7 | 993 | 10.2 | 2312 | 11.5 | Retired | 2526 | 23.1 | 2261 | 23 | 4787 | 23.1 |
| 551-750 | 1769 | 17.0 | 1803 | 18.5 | 3572 | 17.7 | Student/military service | 43 | 0.4 | 17 | 0.2 | 60 | 0.3 |
| 751-1000 | 2271 | 21.8 | 2470 | 25.3 | 4741 | 23.5 | Disability benefit recipient/work assessment allowment | 1239 | 11.4 | 662 | 6.7 | 1901 | 9.2 |
| >1000 | 2257 | 21.7 | 2758 | 28.2 | 5015 | 24.8 | Family income supplement | 7 | 0.1 | 18 | 0.2 | 25 | 0.1 |
| Occupation |  |  |  |  |  |  | Education, mother |  |  |  |  |  |  |
| Unskilled | 1567 | 14.6 | 3018 | 31.1 | 4585 | 22.5 | Primary/lower secondary | 7865 | 73.6 | 7075 | 72.9 | 14,940 | 73.3 |
| Intermediary | 4220 | 39.4 | 1716 | 17.7 | 5936 | 29.1 | Vocational/upper secondary | 1729 | 16.2 | 1671 | 17.2 | 3400 | 16.7 |
| Lower profession | 1486 | 13.9 | 1783 | 18.4 | 3269 | 16.0 | College/university $<4$ years | 697 | 6.5 | 659 | 6.8 | 1356 | 6.7 |
| Higher profession | 3429 | 32.0 | 3173 | 32.7 | 6602 | 32.4 | College/university $>4$ years | 400 | 3.7 | 294 | 3 | 694 | 3.4 |
| NCD |  |  |  |  |  |  | Education, father |  |  |  |  |  |  |
| Diabetes: Yes/yes, previously | 527 | 4.9 | 597 | 6.1 | 1124 | 5.5 | Primary/lower secondary | 6350 | 59.9 | 5690 | 59.1 | 12,040 | 59.5 |
| Heart attack: Yes/yes, previously | 174 | 1.6 | 579 | 6.0 | 753 | 3.7 | Vocational/upper secondary | 2506 | 23.7 | 2400 | 24.9 | 4906 | 24.3 |
| Cancer: Yes/yes, previously | 868 | 8.1 | 768 | 7.9 | 1636 | 8.0 | College/university $<4$ years | 982 | 9.3 | 885 | 9.2 | 1867 | 9.2 |
| Stroke: Yes/yes, previously | 216 | 2.0 | 331 | 3.4 | 547 | 2.7 | College/university $>4$ years | 755 | 7.1 | 653 | 6.8 | 1408 | 7 |
| Childhood economy |  |  |  |  |  |  | Economy, self-evaluated |  |  |  |  |  |  |
| Very good | 667 | 6.2 | 530 | 5.4 | 1197 | 5.8 | Very good | 1826 | 16.8 | 1800 | 18.3 | 3626 | 17.5 |
| Good | 7491 | 69.2 | 6659 | 68.2 | 14,150 | 68.7 | Good | 5468 | 50.3 | 5460 | 55.5 | 10,928 | 52.7 |
| Difficult | 2462 | 22.7 | 2410 | 24.7 | 4872 | 23.7 | Average | 3177 | 29.2 | 2249 | 22.9 | 5426 | 26.2 |
| Very difficult | 205 | 1.9 | 159 | 1.6 | 364 | 1.8 | Difficult | 360 | 3.3 | 258 | 2.6 | 618 | 3 |
| Alcohol, frequency |  |  |  |  |  |  | Very difficult | 50 | 0.5 | 69 | 0.7 | 119 | 0.6 |
| Never | 1089 | 9.9 | 604 | 6.1 | 1693 | 8.1 | Occupation status, self-evaluated |  |  |  |  |  |  |
| Monthly/less | 3066 | 27.9 | 2073 | 20.8 | 5139 | 24.5 | Very high social status | 646 | 6 | 789 | 8.1 | 1435 | 7 |
| 2-4 Times/month | 3954 | 36 | 3945 | 39.6 | 7899 | 37.7 | Fairly high social status | 3376 | 31.6 | 3784 | 38.8 | 7160 | 35 |
| 2-3 Times/week | 2347 | 21.3 | 2628 | 26.4 | 4975 | 23.7 | Neither high nor low status | 5862 | 54.9 | 4591 | 47.1 | 10,453 | 51.1 |
| 4+/Week | 537 | 4.9 | 709 | 7.1 | 1246 | 5.9 | Fairly low status | 735 | 6.9 | 535 | 5.5 | 1270 | 6.2 |
| Smoke daily: Yes, now | 1586 | 14.5 | 1318 | 13.3 | 2904 | 13.9 | Very low status | 66 | 0.6 | 57 | 0.6 | 123 | 0.6 |
| Smoke daily: Yes, previously | 4801 | 43.8 | 4449 | 44.8 | 9250 | 44.3 | Live with spouse: Yes | 7403 | 72.3 | 7880 | 81.6 | 15,283 | 76.8 |
| Smoke daily: Never | 4578 | 41.8 | 4155 | 41.9 | 8733 | 41.8 |  |  |  |  |  |  |  |



Figure 1. The train test methodology as applied in this paper. All steps are repeated when hyperparameters are tuned. All steps of the modelling procedure were performed for each outcome model.
functions that are sufficiently complex to fit the data without fitting the underlying noise. Predictions are evaluated by their ability to forecast the outcome for future data inputs [8]. Figure 1 presents an overview of the train test methodology as applied in this paper. The goal of prediction is twofold; in-sample model fit evaluation, and out-of-sample model validation. For each model, the Troms $\varnothing 7$ sample is therefore partitioned into a training set and a test set. The models are produced on each training set and then applied to each test set. Models are validated by their generalisation error; the prediction error of a model when applied to a general population [6], as represented by the test set.

## The random forest algorithm

We approach this concept of prediction using the random forest algorithm. Random forest estimation grows many decision trees, allowing each tree to vote for the most popular class in classification problems [9]. A major advantage of the random forest approach is its ability to examine non-linear functional forms and complex interaction terms among covariates without the analyst having to prespecify a particular functional form or interaction term [19, 20], or the need for variable transformation [21]. We explore these patterns by first conducting a variable importance analysis by retrieving the estimated mean decrease in accuracy (MDA). MDA scores measure the mean decrease in classification performance
when a given variable is excluded from the model [18] after permuting each element of the set of $X j$ predictors, in which $j$ indexes each covariate, over all trees in the forest [19]. Second, we estimate partial dependences for education, household income and occupation. Partial dependence functions represent the effect of a given variable after accounting for the average effects of the other variables [22]. They represent the functional forms of the association between covariates and outcomes.

We apply the algorithm separately for each NCD. Outcomes are imbalanced, with non-NCD outcomes being much more common for all NCD outcomes. Class imbalances must be considered in classification models to avoid naive predictions. We achieve balanced outcomes by randomly selecting a subset of observations (with replacement) from both outcomes in the training set for each decision tree.

Each tree assumes that the population distribution of the given NCD outcome is equal to $50 \%$, questioning the external validity of our sampling method. Therefore, before training the models and presenting the out-of-bag (OOB) error estimate, we hold out a random set equal to $20 \%$ of the data as a test set. Train test partitions are generated for each individual random forest model. Randomly partitioning the sets from the Tromsø7 dataset preserves the original data generating process, even when data are trained on a balanced subsample.

Random forests are estimated using the R package randomForest [23]. Partial dependences were

Table II. Prediction results from random forest training set and test set.

| Training set |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| NCD | OOB | False negative | False positive | Difference | Sample 0 |
| Diabetes | 0.306 | 0.306 | 0.315 | -0.009 | 200 |
| Stroke | 0.289 | 0.287 | 0.349 | -0.062 | 200 |
| Heart attack | 0.256 | 0.257 | 0.210 | 0.047 | 200 |
| Cancer | 0.340 | 0.343 | 0.312 | 0.030 | 200 |
| Test set |  |  |  |  | 200 |
| NCD | Test error | False negative | False positive | Difference | 200 |
| Diabetes | 0.308 | 0.356 | 0.332 | 0.024 | 400 |
| Stroke | 0.282 | 0.258 | 0.270 | -0.012 | 4001 |
| Heart attack | 0.251 | 0.222 | 0.236 | -0.014 | 4097 |
| Cancer | 0.354 | 0.328 | 0.341 | 4064 | 3897 |

NCD: non-communicable disease; OOB: out-of-bag.
calculated using the pdp package [24]. All models were estimated on 1000 decision trees.

## Results

## Prediction

Table II shows predictive accuracies from both training and test sets. The OOB error rate ranges from $25.6 \%$ for predicting a history of heart attack to $34 \%$ for the cancer outcome. False positives are more common than false negatives, except in diabetes predictions. The difference in classification error rate between classes is small for diabetes ( -0.009 ), slightly higher for cancer (0.03), and the largest for heart attack ( 0.047 ) and stroke ( -0.062 ). Comparing predictions from the test and training sets, we find only minor differences in performance between the OOB and test errors. Error rates for individual classes are also comparable to predictions based on the training set. Predictions from the training set assumed that group sizes were equal, as each tree was fit on a balanced class distribution by undersampling the majority class. Congruence between the test and OOB errors shows that the models translate to unbalanced out-of-sample class distributions, if the algorithm is trained on balanced data.

## Variable importance and partial dependence

Figure 2 presents variable importance in terms of the MDA score for all predictors over all outcomes. The MDA compares the differences between the error rate before and after permuting each predictor variable, normalising the score by dividing the average difference over all trees by their standard deviation [23]. Scores are therefore relative to their variance across all trees in the forest. While we concentrate our presentation on the socioeconomic predictors, variable importance scores are available for all predictors included in the model.


Figure 2. Variable importance (mean decrease in accuracy (MDA) for all predictors in the model, by non-communicable diseases (NCD).

Education performs worst for predicting heart attack, contributes slightly more to predicting cancer and stroke, but is clearly the most effective at predicting diabetes out of the four NCDs. Household income increases the predictive accuracy for cancer, stroke and diabetes, but does not increase predictive accuracy for heart attack. This predictor contributes most to predicting diabetes and stroke, and shows some increase in accuracy in predicting cancer. Occupation increases predictive accuracy for diabetes and stroke outcomes, but seems less effective as a predictor for cancer and heart attack.

Partial dependence plots for education, household income and occupation are presented in Figure 3. Education gradients show decreases in the relative probability of positive classification for all NCD outcomes as education increases, except cancer. For the


Figure 3. Partial dependence plot for three individual socioeconomic indicators. $Y$-axis represents the partial dependence; the proportion of trees voting for positive outcomes, averaged over all predictors in the model. $Y$-axis scaling is not standardised.
cancer outcome, the partial dependence increases post upper secondary education (category 2). For all other NCD outcomes, the education gradient shows minor deviation from a negative linear function.

Overall, partial dependence decreases as household income increases. For the cancer outcome, probabilities increase for those between the lowest and second lowest earners, starts dropping until reaching middleincome earners, showing a flatter but slightly positive gradient among high earners. Partial dependences between diabetes and income show a complex function that broadly separate low, middle and high-income groups, and show minor variations within these broader categories. The relationship between household income and stroke shows a gradient with a sharp decline towards the higher end of the income distribution. For the heart attack outcome, we observe a negative curvilinear reduction in partial dependence, with the effect tapering ever so slightly towards the top of the income distribution. While some minor differences can be found between occupational groups, there are no clear occupational gradients in diabetes and stroke based on prediction. There are greater differences between
unskilled and other occupational groups for the heart attack outcome. The occupational gradient is largely similar to the educational gradient in cancer, likely due to some overlap in the definition of tertiary education and occupational categories 3 and 4 .

## Discussion

We show that education, household income and occupation increase predictive accuracy for several NCDs. The empirical differences between groups are highlighted, as the partial dependences show that the probability of a given individual being classified with a positive outcome varies between most levels of education and income, and to a lesser extent occupation. The relationship between educational attainment, morbidity and mortality is empirically well established $[1,4]$. For the predicted gradient between cancer and education, it is important to note that the outcome measure does not differentiate between cancer types. Certain types of cancer do show reverse SES gradients [1], and this is a likely driver of the results found in this study. However, the reverse gradient is small and defines a difference between participants with upper secondary education and those with tertiary education. Future predictive modelling studies should aim to distinguish different cancer outcomes when calculating gradients.

Household income predictions show a complex picture of NCD prevalence between reported income levels. The income-health gradient is long established, but the precise mechanisms remain up for debate [2, 16]. Results from this study show a complex function with several changes over the income range and substantial variation between the NCDs under study. While we cannot comment on the causal pathways between household income and NCD prevalence, the functional relationship presented by Figure 3 suggests that the issue is not restricted to issues of poverty, given the shape of the income gradients present in diabetes, stroke, heart attack, and to a lesser extent, cancer.

Occupational groups are associated with complex multimorbidities [5]. Occupation and education only partly explain the income-health gradient in Europe [25], suggesting that occupation provides separate causal pathways to health independent from income. A study using the same occupational indicator employed in this study identified an occupational gradient in self-rated health [12]. While we find that occupation improves prediction for diabetes and stroke, occupational gradients in these outcomes are comparatively flat from the perspective of the model. Occupational group differences may be statistically significant [5, 12], but lack predictive capabilities. An
important finding regarding an occupation-health gradient is therefore that statistical significance does not imply accurate prediction, and that occupational gradients are sensitive to model and indicator selection. This is consistent with the general point highlighted in the literature [19] that statistical significance is neither necessary nor sufficient for predictive validity.

Elstad et al. [26] call for future studies on health inequalities in the Nordic countries to embrace causal inference techniques. Following examples in the broader methodological literature [6, 8, 27, 28], we argue that the methodological agenda should be expanded. Algorithmic approaches extend opportunities for causal inference via heterogeneous treatment effect estimation [8]. Further, these methodological paradigms overlap in key areas. For instance, matching techniques in causal inference and ensemble methods such as the random forest algorithm share similar goals in reducing model dependence.

## Strengths and limitations

Our approach has several strengths. The exploratory algorithmic approach to data analysis allows us to highlight the complex functions necessary for predicting a history of NCDs at the individual level. Our results further highlight the importance of indicator selection, and the non-parametric estimation procedure reduces issues that may arise from model dependence. This is important in the context of NCDs as their disease aetiology often defies the expectation of single cause explanations; rather, the risk of a given individual developing these diseases relies on many risk factors. Lundberg [29] argues that the communicative efficiency of paradigms such as the social determinants of health perspective suffers from determinism, and that social regularities cannot be translated into individual predictions because there are large individual variations within social groups. This is an important observation, but in its generalised state it risks missing the forest for the trees. Predictive metrics are necessary precisely because of subpopulation heterogeneity in the individual development of ill health and disease. They highlight not only the extent to which models make correct predictions; they equally highlight the cases in which predictions are wrong and, importantly, how incorrect predictions might occur (e.g. false positives/negatives). Presenting evidence from predictive metrics is therefore to a greater degree congruent with the disease aetiology of the NCDs under study than studies that emphasise statistical significance. Another strength of our study lies in the out-of-sample model validation. The relative congruence between the training and test set predictions shows
that, if the data generating process within the Tromsø7 sample carries over to other data contexts, similar predictions are expected. A clear direction for future studies is therefore to evaluate empirically the extent to which the data generating process in the Tromsø study indeed translates to external data contexts. Our study provides a baseline for which future studies may refine their predictive efforts in the context of health inequalities in NCD prevalence.

Our approach is, however, not without limitations. Simultaneity issues are present in all cross-sectional studies. Interpreting variable importance scores for specific features thus requires some caution. Due to the retrospective and cross-sectional nature of the data, we cannot disentangle those who change their behaviours after receiving a diagnosis from those who do not. This may negatively impact the predictive power of factors such as smoking, alcohol consumption and physical activity. Those individuals with a high healthcare uptake will include those that visit their general practitioner and seek specialist care because of their illness, possibly inflating their predictive importance. The large MDA feature importance score observed for employment status may in part reflect the impact of ill health and health shocks on the probability of employment and labour market exit identified in the literature [30]. Despite these limitations, the MDA scores for education, income and occupational group show that predictive models including socioeconomic indicators will outperform predictive models in which socioeconomic indicators are absent; even in a predictor space including proximal risk factors and indicators sensitive to simultaneity issues. Congruence in predictions between the training set and the test set suggests that overfitting is not a substantial issue when these predictors are included. Further, the partial dependence metrics show that the algorithm uses the information in the socioeconomic indicators to create socioeconomic gradients in health when averaged over all other predictors in the model.

## Conclusions

Results from algorithmic modelling show that the extent to which socioeconomic status contributes to predicting binary NCD outcomes depends on the NCD and the choice of socioeconomic indicator. Evaluating partial dependences reveals that social gradients in NCD outcomes vary in shape between combinations of NCD outcome and socioeconomic indicator. Misclassification rates highlight the extent of variation within socioeconomic groups, suggesting that future studies may improve predictive accuracy by exploring further subpopulation heterogeneity.

There is ample opportunity to leverage predictive modelling further in Norway, due to the vast amount of population data stored in central registries and large cohort surveys. Future studies should apply predictive algorithms in a longitudinal context, such that information on changes in individual behaviour and timing of disease onset can be exploited in investigations on the predictive contributions that socioeconomic status makes.

## Declaration of conflicting interests

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/ or publication of this article.

## Funding

The authors disclosed receipt of the following financial support for the research, authorship and/or publication of this article: This research project was funded by UiT the Arctic University of Norway.

## ORCID iDs

Sigbjørn Svalestuen (iD https://orcid.org/0000-0002-4775-311X
Emre Sari (D) https://orcid.org/0000-0001-9805-9227
Petja Lyn Langholz (i) https://orcid.org/0000-0001-8225-1345
Chi Quynh Vo (iD https://orcid.org/0000-0003-39885726

## References

[1] Mackenbach JP. Health inequalities: persistence and change in modern welfare states. USA: Oxford University Press, 2019.
[2] Kinge JM, Modalsli JH, Øverland S, et al. Association of household income with life expectancy and cause-specific mortality in Norway, 2005-2015. ЭAMA 2019;321:1916-1925.
[3] Balaj M, Huijts T, McNamara CL, et al. Non-communicable diseases and the social determinants of health in the Nordic countries: findings from the European Social Survey (2014) special module on the social determinants of health. Scand $\mathcal{F}$ Public Health 2017;45:90-102.
[4] Lago-Peñas S, et al. The impact of socioeconomic position on non-communicable diseases: what do we know about it? Perspect Public Health 2021;141:158-176.
[5] Vinjerui KH, Bjerkeset O, Bjorngaard JH, et al. Socioeconomic inequalities in the prevalence of complex multimorbidity in a Norwegian population: findings from the cross-sectional HUNT study. BMF Open 2020;10:e036851.
[6] Cranmer SJ and Desmarais BA. What can we learn from predictive modeling? Polit Anal 2017;25:145-166.
[7] McShane BB, Gal D, Gelman A, et al. Abandon statistical significance. Am Stat 2019;73:235-245.
[8] Molina M and Garip F. Machine learning for sociology. Annu Rev Sociol 2019;45:27-45.
[9] Breiman L. Random forests. Mach Learn 2001;45:5-32.
[10] Hopstock LA, Grimsgaard S, Johansen H, et al. The seventh survey of the Tromsø Study (Tromsø7) 2015-2016: study design, data collection, attendance, and prevalence of risk factors and disease in a multipurpose populationbased health survey. Scand $\mathcal{F}$ Public Health 2022;50:919929.
[11] Vos T, Lim SS, Abbafati C, et al. Global burden of 369 diseases and injuries in 204 countries and territories, 19902019: a systematic analysis for the Global Burden of Disease Study 2019. Lancet 2020;396:1204-1222.
[12] Olsen JA, Lindberg MH and Lamu AN. Health and wellbeing in Norway: population norms and the social gradient. Soc Sci Med 2020;259:113155.
[13] Dalene KE, et al. Clustering and trajectories of key noncommunicable disease risk factors in orway: the NCDNOR project. Sci Rep 2023;13:14479.
[14] Brekke KR, et al. Socio-economic status and physicians’ treatment decisions. Health Econ 2018;27:e77-e89.
[15] Vikum E, Krokstad S and Westin S. Socioeconomic inequalities in health care utilisation in Norway: the populationbased HUNT3 survey. Int $\mathcal{F}$ Equity Health 2012;11:1-9.
[16] Pickett KE and Wilkinson RG. Income inequality and health: a causal review. Soc Sci Med 2015;128:316-326.
[17] Breiman L, et al. Statistical modeling: the two cultures (with comments and a rejoinder by the author). Stat Sci 2001;16:199-231.
[18] James G, Witten D, Hastie T, et al. An introduction to statistical learning. New York: Springer, 2013.
[19] Hill DW and Jones ZM. An empirical evaluation of explanations for state repression. Am Pol Sci Rev 2014;108: 661-687.
[20] Jones ZM and Lupu Y. Is there more violence in the middle? Am $\mathcal{F}$ Pol Sci 2018;62:652-667.
[21] Kreatsoulas C and Subramanian S. Machine learning in social epidemiology: learning from experience. SSM Popul Health 2018;4:347.
[22] Friedman JH and Meulman JJ. Multiple additive regression trees with application in epidemiology. Stat Med 2003;22:1365-1381.
[23] Liaw A and Wiener M. Classification and regression by randomForest. R News 2002;2:18-22.
[24] Greenwell BM. Pdp: An R package for constructing partial dependence plots. $R \mathscr{F}$ 2017;9:421-436.
[25] Huijts T, Eikemo TA and Skalická V. Income-related health inequalities in the Nordic countries: examining the role of education, occupational class, and age. Soc Sci Med 2010;71:1964-1972.
[26] Elstad JI, Heggebø K and Dahl E. Nordic research on health inequalities: a scoping review of empirical studies published in Scandinavian fournal of Public Health 2000-2021. Scand F Public Health 2022;50:843-851.
[27] Seligman B, Tuljapurkar S and Rehkopf D. Machine learning approaches to the social determinants of health in the health and retirement study. SSM Popul Health 2018;4:95-99.
[28] Broderstad TS. An empirical evaluation of explanations for political system support. Polit Res $Q$ 2023; 76: 10659129231156388.
[29] Lundberg O. Next steps in the development of the social determinants of health approach: the need for a new narrative. Scand $\mathcal{F}$ Public Health 2020;48:473-479.
[30] García-Gómez P. Institutions, health shocks and labour market outcomes across Europe. $\mathcal{F}$ Health Econ 2011;30:200-213.


[^0]:    Correspondence: Sigbjørn Svalestuen, NORCE Norwegian Research Centre AS, Siva Innovasjonssenter, Sykehusvn 21, 9019 Tromsø, Norway. E-mail: ssva@norceresearch.no

