1	<u>Title</u>		
2	Marine heatwaves are not a dominant driver of change in demersal fishes		
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4			
5	Authors		
6	Alexa L. Fredston, William W. L. Cheung, Thomas L. Frölicher, Zoë J. Kitchel, Aurore A.		
7	Maureaud, James T. Thorson, Arnaud Auber, Bastien Mérigot, Juliano Palacios-Abrantes, Maria		
8	Lourdes D. Palomares, Laurène Pecuchet, Nancy Shackell, Malin L. Pinsky		
9			
10	1. Alexa L. Fredston * <i>corresponding author</i>		
11	fredston@ucsc.edu		
12	Department of Ocean Sciences, University of California, Santa Cruz, Santa Cruz,		
13	California, USA		
14	2. William W. L. Cheung		
15	w.cheung@fisheries.ubc.ca		
16	Institute for the Oceans and Fisheries, University of British Columbia, Vancouver, British		
17	Columbia, Canada		
18	3. Thomas L. Frölicher		
19	thomas.froelicher@unibe.ch		
20	Climate and Environmental Physics, Physics Institute, University of Bern, Bern,		
21	Switzerland		
22	Oeschger Centre for Climate Change Research, University of Bern, Bern, Switzerland		

23	4.	Zoë J. Kitchel
24		zoe.kitchel@rutgers.edu
25		Department of Ecology, Evolution, and Natural Resources, Rutgers University, New
26		Brunswick, New Jersey, USA
27	5.	Aurore A. Maureaud
28		aurore.aqua@gmail.com
29		Department of Ecology and Evolutionary Biology, Yale University, New Haven,
30		Connecticut, USA
31		Department of Ecology, Evolution, and Natural Resources, Rutgers University, New
32		Brunswick, New Jersey, USA
33	6.	James T. Thorson
34		james.thorson@noaa.gov
35		Alaska Fisheries Science Center, National Marine Fisheries Service, National Oceanic
36		and Atmospheric Administration, Seattle, Washington, USA
37	7.	Arnaud Auber
38		Arnaud.Auber@ifremer.fr
39		Institut Français de Recherche pour l'Exploitation de la MER (IFREMER), Unité
40		Halieutique Manche Mer du Nord, Laboratoire Ressources Halieutiques, Boulogne-sur-
41		Mer, France
42	8.	Bastien Mérigot
43		bastien.merigot@umontpellier.fr

44 MARBEC, University of Montpellier, CNRS, Ifremer, IRD, Sète, France

45	9.	Juliano Palacios-Abrantes
46		j.palacios@oceans.ubc.ca
47		Institute for the Oceans and Fisheries, University of British Columbia, Vancouver, British
48		Columbia, Canada
49	10	. Maria Lourdes D. Palomares
50		m.palomares@oceans.ubc.ca
51		Sea Around Us, Institute for the Oceans and Fisheries, University of British Columbia,
52		Vancouver, British Columbia, Canada
53	11	. Laurène Pecuchet
54		laurene.pecuchet@uit.no
55		The Arctic University of Norway, Tromsø, Norway
56	12	. Nancy Shackell
57		Nancy.Shackell@dfo-mpo.gc.ca
58		Bedford Institute of Oceanography, Fisheries and Oceans Canada, Dartmouth, Canada
59	13	. Malin L. Pinsky
60		malin.pinsky@rutgers.edu
61		Department of Ecology, Evolution, and Natural Resources, Rutgers University, New
62		Brunswick, New Jersey, USA
63		Department of Ecology and Evolutionary Biology, University of California, Santa Cruz,
64		Santa Cruz, California, USA
65		
66	<u>Sumn</u>	<u>nary</u>

Marine heatwaves have been linked to negative ecological effects in recent decades.<sup>1,2</sup> If marine 68 69 heatwaves regularly induce community reorganization and biomass collapses in fishes, the consequences could be catastrophic for ecosystems, fisheries, and human communities.<sup>3,4</sup> 70 71 However, the extent to which marine heatwaves have negative impacts on fish biomass or 72 community composition or even whether their effects can be distinguished from natural and 73 sampling variability remains unclear. We investigated the effects of 248 sea bottom heatwaves from 1993 to 2019 on marine fishes by analyzing 82,322 hauls (samples) from long-term 74 scientific surveys of continental shelf ecosystems in North America and Europe spanning the 75 76 subtropics to the Arctic. We found that the effects of marine heatwaves on fish biomass were 77 often minimal and could not be distinguished from natural and sampling variability. Further, 78 marine heatwaves were not consistently associated with tropicalization (gain of warm-affiliated 79 species) or deborealization (loss of cold-affiliated species) in these ecosystems. While steep declines in biomass occasionally occurred after marine heatwaves, these were the exception, not 80 the rule. Against the highly variable backdrop of ocean ecosystems, marine heatwaves have not 81 82 driven biomass change or community turnover in fish communities that support many of the world's largest and most productive fisheries. 83

84

#### 85 <u>Main text</u>

86

Extreme climatic events exacerbated by global climate change are associated with many
examples of ecological transformation<sup>5</sup>. Marine heatwaves (MHWs)<sup>6</sup>—prolonged periods of
anomalously warm ocean temperatures—have been linked to widespread coral bleaching and
die-offs of kelp forests and reef fishes in shallow coastal seas<sup>1,2</sup>. MHWs can rapidly displace

91 some marine species by hundreds of kilometers and cause abrupt declines in phytoplankton and commercially important species<sup>7-9</sup>. These high-profile catastrophes and those emerging from 92 regional and global model simulations predict that MHWs will likely wreak ecological 93 devastation and negatively impact socio-economic systems<sup>3,4,10</sup>. However, accurate predictions 94 must rely upon generalizable patterns and processes, not case-studies with limited spatial and 95 96 taxonomic scope. The extent to which MHWs in general have negative ecological impacts or 97 even whether they can be differentiated from other sources of natural and sampling variability in marine systems remains unclear. In the few studies that have compared responses to MHWs 98 across multiple species within the same ecosystem, some species declined while others thrived<sup>11-</sup> 99 <sup>13</sup>, suggesting that single-species responses do not accurately reveal net ecological effects. These 100 101 net effects are particularly important to understand in continental shelf ecosystems, where many of the world's largest ocean fisheries operate<sup>14</sup>. 102

103 Here, we analyzed the cross-species and cross-ecosystem effects of MHWs in 104 Northern-hemisphere shelf ecosystems from the subtropics to the Arctic. We analyzed 82,322 105 hauls (discrete samples) comprising 22,574,452 observations of 1,769 demersal fish taxa from 18 106 long-term scientific (i.e., fisheries-independent) bottom trawl surveys covering 45 degrees of 107 latitude in the Northeast Pacific, Northwest Atlantic, and Northeast Atlantic (Fig. 1). Ninety 108 percent of hauls occurred in waters 18-448 m deep. These surveys represent the most spatially, temporally, and taxonomically extensive observations available for testing MHW effects. We 109 110 examined observations from 1993-2019 to quantify the effects of MHWs on regional fish 111 biomass and community composition (Fig. 1). We defined a MHW as a period of  $\geq$  5 days with sea bottom temperature (SBT) 112

anomalies above the seasonally varying 95<sup>th</sup> percentile for that region<sup>6</sup> and used cumulative

114 intensity as the primary metric characterizing MHWs (see Methods). Cumulative intensity is 115 measured in °C-days and represents the sum of the time-varying anomalies above the 95<sup>th</sup> 116 percentile threshold over the duration of the event. This approach—defining MHWs as physical phenomena based on relative temperature anomalies—is widely applied in oceanography.<sup>15</sup> A 117 contrasting approach is to define absolute temperature thresholds above which deleterious 118 ecological impacts consistently occur, as have been identified for coral reefs.<sup>16</sup> Such a biological 119 120 threshold has not been described in most marine systems, including the ones we studied, so we 121 explored a range of relative and absolute MHW metrics (see Methods). We hypothesized that 122 MHWs altered fish biomass and community composition and that these effects would increase 123 with the cumulative intensity of MHWs.

This dataset recorded some notable MHW impacts that mirror previous reports in the 124 125 literature, including a 22% biomass loss in the Gulf of Alaska following the 2014-2016 Northeast Pacific MHW with a cumulative intensity of 57 °C-days<sup>8,17</sup> and a 70% biomass gain in 126 the Northeast USA following the 2012 Northwest Atlantic MHW (67 °C-days; Fig. 2a, 3a)<sup>11</sup>. 127 128 However, it is important to note that while these effects were substantial, they were neither large 129 compared to natural variability in biomass nor repeated across other previously unreported 130 MHWs. Other intense MHWs had little discernible effect on total biomass, such as the 42 °Cdays MHW preceding the 2008 North Sea survey that recorded only a 6% biomass decline (Fig. 131 132 2a). Some of the largest biomass changes occurred in non-MHW years, such as the 97% increase in biomass in the North Sea in 2011, or the 77% biomass decline in the Southeast US in 1996.<sup>18</sup> 133 In addition, we observed that the most extreme biomass changes were often reversed in 134 subsequent years. For example, the southern Gulf of St. Lawrence survey hauls caught an 135 average of 4 metric tons of fish per km<sup>2</sup> in 2011, a non-MHW year; 13 metric tons per km<sup>2</sup> in 136

2012 following 36 °C-days MHW; and 4 metric tons per km<sup>2</sup> in 2013 following a MHW with 0.4
°C-days of cumulative intensity (Fig. 2a). The peak in biomass in 2012 was driven by Atlantic
herring (*Clupea harengus*) and likely reflects a combination of survey variability and a true peak
in population size of the fall Atlantic herring stock in the region<sup>19</sup>.

- 141 Contrary to our expectations, both among surveys that were and were not preceded by a
- 142 MHW, the median change in biomass was close to zero  $(0.023 \pm 0.367 \text{ and } 0.016 \pm 0.323,$

143 respectively; medians and standard deviations of log ratios), indicating that regions were

approximately as likely to exhibit net biomass gains as losses (Fig. 2a). Of the 369 survey-years

145 we matched with sea bottom temperature data, 139 followed MHWs and 230 did not. There was

146 no significant difference in mean biomass change between surveys that were and were not

147 preceded by a MHW (p = 0.40, t = 0.85, df = 249; two-sided t-test; Fig. 2a). Further, the

148 cumulative intensity of a MHW had no significant relationship with the change in biomass

149 (linear regression;  $R^2 = 0.00$ ; p=0.88; Fig. 2, Supp. Tab. 2).

150 The ecosystems we studied have distinct climates, species assemblages, and histories of anthropogenic pressures<sup>20</sup>, and might respond at different rates and in different directions to 151 environmental perturbation.<sup>21</sup> More broadly, pulse disturbances and other exogenous drivers 152 153 (including heatwaves) are often expected to increase variance in the biomass of populations and 154 communities<sup>22</sup>. However, we find no statistically significant relationship between the cumulative 155 intensity of a MHW and biomass change of these demersal fish assemblages in any individual 156 region (Extended Data Fig. 1, Supp. Tab. 3). Across all surveys, variability-measured as the 157 absolute value of the year-over-year biomass log ratios-did not increase with MHW cumulative intensity (linear regression,  $R^2 = 0.00$ , p = 0.24; Fig. 2b, Supp. Tab. 4). Instead, we find that 158 159 variability in biomass change from one year to the next is similarly high with or without MHWs

160  $(0.22 \pm 0.248 \text{ and } 0.19 \pm 0.214, \text{ respectively; medians and standard deviations of absolute log}$ 161 ratios) and that these absolute log ratios of biomass are not significantly different (p = 0.24, t = 162 1.17, df = 245; two-sided t-test; Fig. 2b).

163 Accounting for latitude, depth, temporal lags, autoregression in the biomass time-series, 164 fisheries catch, and species traits (feeding mode, trophic level, and habitat) also did not reveal 165 any meaningful effect sizes of MHWs on biomass (Supp. Tab. 6-10, Extended Data Fig. 6-7). 166 Our results were also robust to the metric used (cumulative intensity, duration, intensity, or degree heating days) to characterize MHWs, how cumulative intensity was scaled, and to 167 168 whether SBT data were detrended (Extended Data Fig. 2, Supp. Tab. 5). Because SBT data was 169 only available from 1993 onward, we also analyzed a longer time-series of sea surface 170 temperature (SST) that began in 1982 and thus could be paired additional fish surveys. The SST 171 analysis included 100,877 hauls comprising 26,886,245 discrete taxon observations, and yielded results that were qualitatively similar to the SBT results described in the main text (Extended 172 Data Fig. 2b). Because deleterious heatwave effects have often been recorded in summer<sup>1</sup>, we 173 174 also tested for an effect of summer-only MHWs on biomass (Extended Data Fig. 2g), finding a weak positive effect (i.e., greater biomass following more intense MHWs; linear regression,  $R^2 =$ 175 0.02, p = 0.02). While interpreting this result cautiously given the high leverage of a few data 176 177 points, this result is consistent with the 2012 Northwest Atlantic MHW that occurred in summer and was associated with an increase in biomass in numerous fisheries<sup>11</sup>. 178

Individual MHWs may lead to rapid ecological turnover by causing cold-affiliated
species to decline or go extinct ("deborealization") and/or by causing warm-affiliated species to
spread or increase ("tropicalization")<sup>23-25</sup>. We tested whether tropicalization or deborealization
are general effects of MHWs by calculating the Community Temperature Index (CTI) for each

survey in each year and comparing CTI change to MHW occurrence and cumulative intensity
(Methods). CTI is an aggregate thermal niche index for the entire community calculated as the
biomass-weighted mean of single-species' realized thermal niches<sup>26</sup>.

186 Other studies show that CTI has increased in North American fish communities in recent decades, concomitant with ocean warming.<sup>25</sup> To explore whether MHWs induce CTI increases, 187 we first focused on the 2014-2016 Northeast Pacific MHW, nicknamed "The Blob"-one of the 188 189 largest MHWs in our dataset (Fig. 2, 3). Of the four regions in the Northeast Pacific for which 190 we had data, the Gulf of Alaska exhibited the most pronounced CTI increase after the 2014-2016 MHW—from 7.25 °C in 2013 to 7.39 °C in 2015 and 7.50 °C in 2017. We also found a CTI 191 192 increase in the West Coast USA in 2015 following an 8 °C-days SBT MHW (CTI change 0.11 °C), consistent with findings that warm-affiliated subtidal fishes increased in Southern California 193 that year<sup>27</sup>. However, CTI in the Eastern Bering Sea decreased by 0.22 °C from 2015 to 2017. In 194 195 British Columbia, we measured CTI values of 8.34 °C in 2013, 8.10 °C in 2015, and 8.31 °C in 2017 (Fig. 3b). 196

Our analysis found no evidence for systematic tropicalization or deborealization in marine fish communities across all 18 surveys and 369 survey-years in response to MHWs (Fig. 4). Year-over-year CTI change in communities that did not experience MHWs was not significantly different from those that did  $(0.024 \pm 0.996 \text{ and } 0.007 \pm 0.983 \text{ °C}, \text{ respectively};$ means and standard deviations; p = 0.87, t = 0.16, df = 280; two-sided t-test; Fig. 4b). Further, there was no relationship between MHW cumulative intensity and CTI change ( $\mathbb{R}^2 = 0.00$ ;  $\mathbf{p} = 0.33$ ; linear regression; Supp. Tab. 11).

Marine heatwaves may restructure ecological communities in other ways beyond
 tropicalization and deborealization<sup>28</sup>. Less predictable changes in species identity and underlying

community structure could also emerge despite consistent biomass<sup>29</sup>. We tested for changes in 206 207 community composition by calculating dissimilarity over the time-series using occurrence data 208 (i.e., species presence-absence) as well as biomass data. Each of these dissimilarity indices is 209 calculated from one year to the next using two components-one measure of how much species 210 are substituted for one another and one measure of how much each community is a subset of the other<sup>30</sup> (see Methods). A high dissimilarity value between two years may be driven by a 211 212 transition away from the baseline community structure as a result of disturbance, which has been 213 observed in marine systems in response to climate change<sup>29</sup>.

214 In some instances, fish communities exhibited high dissimilarity from the previous year 215 following a MHW, such as in the Eastern Bering Sea and the West Coast US during the 2014-216 2016 MHW (Fig. 3c). However, this was not a general effect. We found that community 217 dissimilarity measured between consecutive years was not, on average, significantly different 218 whether or not a MHW occurred when measured with occurrence-based substitution (p = 0.12, t = 1.57), biomass-weighted substitution (p = 0.99, t = -0.02), or biomass-weighted subset (p =219 220 0.32, t = 1.00; all two-sided t-tests; Extended Data Fig. 8). The one statistically significant 221 relationship suggested that the subset component of occurrence-based dissimilarity was smaller 222 after MHW years than after non-MHW years (p = 0.01, t = -2.52, two-sided t-test; Extended 223 Data Fig. 8b), the opposite of the hypothesized effect. This observed community stability in the face of MHWs could indicate that climate refugia, such as depth refugia or other thermal refugia, 224 provide safe havens for species during extreme events.<sup>31</sup> Further, changes in community 225 226 structure at the local scale may not be reflected at the regional scale of our analysis. 227 This array of results suggests that the regional impacts of MHWs on fish communities are 228 highly idiosyncratic, with dramatic effects in single cases but not in general. In particular, the

229 effects of MHWs do not yet exceed natural variability in these ecosystems, or the variability due 230 to the sampling process. These results also highlight the need to further explore contextdependent responses<sup>32</sup>. Species and spatial portfolio effects<sup>33,34</sup> as well as spatial and temporal 231 storage effects<sup>35</sup> may explain individualistic responses to extreme events that could buffer many 232 ecosystems from MHWs. Range shifts and mortality and fecundity rates may vary with 233 latitude,<sup>25</sup> although we did not find a latitude effect in our analysis (Supp. Tab. 8). Interspecific 234 variation in the timing, magnitude, and direction of MHW response may also be important<sup>36</sup>. For 235 236 example, population dynamics of abundant species in response to the environment and fishing 237 drive some of the biomass changes we observed (Extended Data Fig. 10). Ecological responses 238 to warming may also be mediated by direct and indirect effects of other human impacts on the oceans, such as fishing, fisheries management, and changing primary productivity<sup>3</sup>. 239

240 To verify that our dataset had sufficient statistical power, we developed a series of power 241 analyses. First, we estimated that our dataset (n = 369 survey-years paired with SBT data) had 242 the power to reveal a consistent MHW-induced regional fish biomass decline of 9% or greater 243 (Extended Data Fig. 9c). Using the longer time-series of 441 survey-years that we paired with SST data had the power to detect a biomass decline of 8% or greater (Extended Data Fig. 9d). 244 Such an effect did not emerge from the results of this study, suggesting that any MHW effects 245 246 are smaller than this. We note that even decline in fish biomass of 8-9% or less, if permanent and 247 sustained over time, would likely have substantial deleterious consequences for marine fisheries and social-ecological systems<sup>3</sup>. One model simulation of marine fishes experiencing MHWs 248 249 under the high emission, no mitigation future climate scenario (RCP 8.5) projected that the 250 negatively affected stocks (approximately <sup>3</sup>/<sub>4</sub> of total stocks) would exhibit an average biomass decline of 6%<sup>3</sup>. Approximately 600 survey-years would be required to detect an average biomass 251

decline of 6% (Extended Data Fig. 9a, b). Put another way, MHWs may have had effects ondemersal fish communities in recent decades, but if so, those effects were small.

- 254 In addition, there are alternative explanations to consider. MHWs may affect survey 255 methods themselves: a study on several coral trout species found that they were more 256 catchable—i.e., encountered by fishing gear at higher rates—in warmer temperatures<sup>37</sup>. 257 However, if this response was widespread among the species we studied, it would cause an 258 increase in biomass following MHWs that we did not observe. The availability of fishes to surveys can also be influenced by range shifts, possibly induced by MHWs<sup>38</sup>. However, our 259 260 analysis accounted for this by testing for tropicalization or deborealization within fixed spatial 261 areas (the survey regions). Because biomass trends may be strongly structured by commercial fisheries catches, we fitted models predicting biomass change with fisheries catch as well as 262 263 MHW cumulative intensity, with no significant results (Supp. Tab. 10). While the choice of metric to quantify MHWs<sup>6,15,38,39</sup> and fish community responses<sup>11,37</sup> may influence results, our 264 265 results here were not sensitive to these decisions. The pattern in Fig. 2 emerged regardless of 266 whether we analyzed fish biomass (i.e., weights) or fish abundance (i.e., counts) or whether 267 taxon-level records were summarized as means or medians (Extended Data Fig. 5). The spatial scale of our study was determined by the surveys, which themselves are designed to capture 268 269 distinct biogeographical and political regions and/or to follow fisheries management criteria<sup>40</sup> 270 (see Methods). However, it is possible that fish community responses to MHWs vary with the spatial scale at which they are measured, as has been found with other metrics of biodiversity.<sup>41</sup> 271 Ecological effects of climate change result from the interaction of long-term climatic 272 change combined with short-term extreme events such as heatwaves<sup>5</sup>, which have been projected 273
- to cause widespread ecological devastation on land and in the sea (although see  $^{36,42}$ ). To date,

275 this expectation has largely been based on case studies that select one or a few prominent species 276 and ecosystems with remarkable MHW responses, rather than the comprehensive approach that 277 we used. Analyses that focus on particular species that were historically prevalent in a region 278 may be predisposed to find a biomass decline following a MHW, possibly due to unrelated time-279 series dynamics (e.g., mean reversion or density dependence) or because species that were 280 dominant in historical climates might not be as successful after a MHW. In addition, case study approaches tend to select and emphasize extreme responses<sup>43</sup>—although several regional studies 281 282 focused on a single MHW event have also found no net loss of abundance or biomass in coral- or kelp forest-associated fishes<sup>27,44</sup>. Further empirical research that builds on the present study and 283 284 extends beyond individual case studies is needed to interpret, contextualize, and predict severe MHW effects.<sup>2</sup> Building partnerships to leverage existing non-public datasets from the southern 285 286 hemisphere and other under-studied parts of the global ocean will also be helpful in understanding how ecological context influences MHW responses<sup>40</sup>. 287

288 Our findings highlight the need to understand divergent responses to extreme events. 289 Single-species responses may be mediated by thermal tolerances, but we did not find evidence 290 that cold-affiliated species decline or that warm-affiliated species increase following MHWs (Fig. 4a). Other studies find that species' responses vary from one extreme event to another<sup>42</sup>. 291 292 Portfolio and storage effects may explain why ecosystem-level MHW effects are rare, but they 293 do not reveal what caused certain MHWs to have deleterious ecological effects. The community 294 stability we observed in the face of MHWs could indicate that climate refugia, such as depth or other thermal refugia, provide safe havens for species during extreme events.<sup>31</sup> Cumulative 295 impacts of MHWs and other stressors such as harmful algal blooms<sup>13</sup> or low-productivity events<sup>9</sup> 296 297 could play a role. Perhaps very extreme MHWs in the future will cross a tipping point beyond

298 which adverse ecological effects occur, but we did not see such a tipping point in the recent 299 historical record. Other fields (e.g., coral reef ecology) have identified such thresholds, though the generality of thresholds across ecological systems remains unclear.<sup>45</sup> Gaining mechanistic 300 301 insight into why only some MHWs have deleterious effects, and on only some species, is 302 necessary for any future efforts to identify an effect threshold or forecast MHW impacts and 303 should be a research priority for the field. Additionally, ecosystem responses to extreme "pulse" 304 events such as MHWs can shape impacts of more gradual "press" trends; the complex interactions between these climate change effects warrant future research.<sup>5</sup> 305 306 Understanding MHW impacts on entire ocean ecosystems is particularly crucial in the 307 context of accelerating global change and efforts to advance towards ecosystem-based management that considers the many links between species and with their environment<sup>46</sup>. MHW 308 309 occurrences are projected to emerge above their natural variability within this century in many regions<sup>47</sup>. Future research will be needed to determine the extent to which fish community 310 311 impacts of MHWs will grow as MHWs intensify, or whether portfolio and other ecological 312 effects can buffer ecosystems from MHW impacts. Marine life is more vulnerable to warming 313 than terrestrial life, because marine organisms tend to live close to their thermal limits and fewer thermal refugia exist in the seas<sup>48</sup>. Observed and predicted changes in marine ecosystems in 314 315 response to global warming formed part of the rationale behind the Paris Climate Agreement to limit the global mean surface temperature increase above industrial levels to 1.5 °C by 2100<sup>49</sup>. 316 As a future that is more than 1.5 °C warmer looks increasingly likely<sup>50</sup>, it is more critical than 317 318 ever to develop a deeper understanding of what drives ecological responses to extreme climate 319 events.

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- 431
- 432 <u>Figures</u>



434 Fig. 1. Of 18 regions studied from the Atlantic and Pacific Oceans, all experienced marine heatwaves during the available scientific fish survey time-series. Highlighted areas on the 435 map represent the spatial area surveyed by each trawl survey. Inset plots show the number of 436 437 distinct sampling events (i.e., hauls of the trawl net) every year (grey bars, right y-axis) and the MHW cumulative intensity in °C-days calculated from sea bottom temperature data (lines, left v-438 axis; warmer colors represent greater cumulative intensity). Years correspond to "survey 439 440 reference years"—the twelve months preceding the survey—rather than calendar years, and vary 441 among regions (see Methods).



444 Fig. 2. More intense marine heatwaves were not associated with a decline in fish biomass or an increase in biomass variability, and biomass was approximately as likely to increase as it 445 was to decrease from one year to the next, regardless of whether a marine heatwave 446 447 occurred. MHW cumulative intensity (°C-days) had no relationship with (a) biomass change (measured as log ratio; linear regression,  $R^2 = 0.00$ , p = 0.45) or (b) absolute biomass change 448 (measured as absolute log ratio; linear regression,  $R^2 = 0.01$ , p = 0.07). The grey shaded area is a 449 95% confidence interval. Density plots along the right-hand y-axes show that biomass changes 450 from one year to the next were the same whether a MHW did (red, n = 139) or did not (blue, n =451 452 230) occur. Points represent (a) log ratios or (b) absolute log ratios of mean biomass in a survey from one year to the next (n = 369). Surveys following the highest cumulative intensity MHWs 453 454 are labeled (years correspond to when the survey was conducted). Colors correspond to regions 455 in Fig. 1. Models exploring the relationships in (a) and (b) are reported in Supp. Tab. 2 and 4, 456 respectively.











ecosystem boundaries (e.g., Large Marine Ecosystems<sup>52</sup>) and are similar in size to the spatial 485 extents of MHWs<sup>39</sup>. Although these surveys are conducted by many nations, we used only 486 487 surveys by agencies that make their raw data publicly available, which facilitates reproducibility 488 of this study. Those agencies were the National Oceanic and Atmospheric Administration 489 (NOAA), Fisheries and Oceans Canada (DFO), the International Council for the Exploration of 490 the Sea (ICES), and the Institute of Marine Research (IMR). This limitation constrained our 491 analysis to the Northeast Pacific, Northwest Atlantic, and the Northeast Atlantic (Supp. Tab. 1). In these surveys, the sampling unit is a single haul, i.e., a sampling event in which a net is 492 493 towed through the water. We filtered invalid hauls based on reported sample quality, sampling 494 times, and availability of variables required to calculate taxon-level biomass data. Further detail on data cleaning and harmonization in addition to raw data and code can be found in Maureaud 495 et al<sup>53</sup>. All primary analyses used biomass (weight) data; we conducted a supplementary analysis 496 497 of the main results using abundance (count) data for the regions for which it was available (all 498 but Northeast US; Extended Data Fig. 5). 499 Datasets were trimmed to standardize the spatial footprint of the survey over time, to match the available temperature datasets (GLORYS began in 1993 and OISST began in 1982; 500

see Marine heatwave data), to remove years with very few samples, and to omit samples
collected outside of the focal season (3-month interval) of each survey. We used the World
Register of Marine Species<sup>54</sup> to standardize taxonomies, and the "dggridR" R package<sup>55</sup> to
standardize the survey footprints. Across our 18 surveys, we paired 82,322 hauls with GLORYS
and 101,376 hauls with OISST. Of the 94% of hauls with an associated depth value, 90%
occurred between 18 m and 451 m depth. The depths sampled vary due to each region's unique
bathymetry and each survey's protocols, but surveys tend to sample similar depths over time: for

508 example, the Southeast US survey samples very shallow inshore areas of just 3 m depth, while 509 the Scotian Shelf and West Coast surveys routinely trawl deeper than 1000 m. We used all taxa 510 for biomass analyses but only the species-level observations for community analyses (see

511

# **Species and Community Temperature Indices**).

512 We imputed zeros representing an observed absence in every instance when a species 513 (that was observed at some point in the region) was not recorded in a haul. These absences can 514 be considered true non-detections due to the standardized spatiotemporal design of bottom trawl 515 surveys. We then calculated a mean biomass for each species in every year, and calculated 516 region-wide biomass as the sum of species-level biomass. To assess the sensitivity of our results 517 to the metric used, we also calculated median biomass, mean abundance, and median abundance 518 in the same way (although the Northeast US region did not have abundance data and was thus 519 omitted from the abundance analysis). We did this across the entire survey domain (following 520 the spatial standardization mentioned above), rather than within the "strata" used in some analyses<sup>56</sup>, because not all of the surveys have stratified sampling designs and we wanted to be 521 522 consistent across all regions.

Year-over-year mean biomass change was calculated as a natural log ratio, 523

 $\ln(\frac{biomass_t}{biomass_{t-1}})$ . Log ratios for median biomass, mean abundance, and median abundance were 524 525 calculated the same way. For straightforward interpretation, we also reported percentage biomass 526 changes in the text, although biomass log ratios were used in all models. For example, a 67% biomass increase means that  $biomass_t = 1.67 \times biomass_{t-1}$ . A 67% decrease means that  $biomass_t$ 527  $= (1 - 0.67) \times biomass_{t-1}$ . 528

Because surveys began in different months, we paired each survey's biomass data with 529 MHW data from the preceding 12 months. For example, for a survey that began in August, the 530

531 August 2010 data was paired with MHW data from August 2009 - July 2010. Because we have 532 no a priori information on the season in which MHWs could have the greatest ecological 533 impact—and this may vary by species and life stage—we analyzed MHW effects over a full year (i.e., 12-month interval). Some substantial MHW effects have been reported in summer,<sup>57</sup> and 534 yet winter heatwaves strongly influence distribution and abundance for species limited by winter 535 survival<sup>58,59</sup>. Warmer winters are hypothesized to have driven recent ecological changes in the 536 Gulf of Alaska<sup>12</sup>. Winter MHWs may also reduce recruitment in habitat-forming seaweeds<sup>6</sup> and 537 cause metabolic stress to coral reef fishes<sup>60</sup>. Spawners and embryos have narrower temperature 538 tolerance ranges than non-spawning adults<sup>61</sup>; a spring heatwave could thus affect the survival or 539 540 performance of spring-spawning fishes and their embryos. To test the sensitivity of our results to this choice, we also explored the effects of only summer anomalies on biomass change 541 542 (Extended Data Fig. 2g).

Interannual biomass variability is significantly lower in surveys with more samples per year (linear regression;  $R^2 = 0.08$ ; p < 0.001). To account for this, all models and statistical tests either included a survey effect or used biomass log ratios that were scaled and centered within surveys.

Marine heatwave data. We paired the demersal bottom trawl data with MHWs calculated with sea bottom temperature (SBT) data from the Copernicus 1/12° (about 8 km) global ocean reanalysis, the Global Ocean Reanalysis and Simulations (GLORYS12).<sup>62</sup> The reanalysis dataset is generated with the Nucleus for European Modelling of the Ocean (NEMO) ocean model forced by the ERA-Interim atmospheric reanalysis. The model assimilates satellite altimetry, satellite SST, sea ice concentrations, and in-situ profiles of salinity and temperature. We chose GLORYS12 for our analyses because it provided daily estimates of temperature anomalies at depth at a fine spatial resolution, and it reproduces nearshore bottom temperatures
and recent MHWs with enhanced fidelity compared to other products.<sup>63,64</sup> Being highly
dependent on ocean observations for data assimilation, GLORYS12 only began in 1993. We
used the 1993-2019 period for our analysis.

558 Because many bottom trawl datasets began earlier than 1993, we also calculated SST 559 MHWs to conduct supplementary analyses with a longer time-series. For SST, we used the 560 NOAA daily Optimum Interpolation Sea Surface Temperature (OISST) Analysis version 2.1 dataset<sup>65,66</sup> with a horizontal grid resolution of 0.25°, which is available from 1982 onward, to 561 562 characterize MHWs. This dataset provides a daily global record of surface ocean temperature 563 observations obtained from satellites, ships, buoys, and Argo floats on a regular grid. Infrared satellite data from the Advanced Very High Resolution Radiometer is its main input and any 564 565 large-scale satellite biases relative to in-situ data from ships and buoys are corrected. Gaps are filled in by interpolation. We used the 1982-2019 period for our analysis. OISST and GLORYS 566 are plotted against one another in Extended Data Fig. 4. 567

Following standard MHW definition (e.g. 9,47), both SBT and SST anomalies were 568 569 calculated within (not across) spatial units—here, the survey regions. This approach defined 570 anomalies relative to historical conditions in a region, which are likely reflective of the 571 environments to which organisms are adapted. In other words, each MHW we identified represented a departure from whatever climate the marine organisms in that region typically 572 573 experienced. Because our study regions varied substantially in seasonality, natural variability, 574 and exposure to oceanographic phenomena, calculating anomalies from cross-region pooled SBT 575 and SST values would identify only the most globally extreme events as MHWs and would omit

the many events in which temperatures were anomalously high for species within a region butnot necessarily high for the global oceans.

For both SBT and SST, we defined a MHW as a period of at least five continuous days during which the SBT (or SST) averaged for each survey area was larger than a seasonally varying threshold given by the 95<sup>th</sup> percentile of the survey-area averaged SBT (or SST) anomalies (relative to the mean seasonal cycle that is calculated for each calendar day individually). This is a common approach for defining MHWs, although some authors use the 90<sup>th</sup> percentile<sup>6</sup> or the 99<sup>th 39</sup> instead of the 95<sup>th</sup>. Under our definition, MHWs may occur throughout the year and at all locations.

585 The temperature data was linearly detrended before any analysis to distinguish discrete MHWs from the long-term warming signal (see Jacox et al. <sup>38,67</sup> and Extended Data Fig. 3) 586 587 although we also tested the sensitivity of our results to this decision by re-running the analysis 588 with non-detrended data and reached equivalent conclusions (see Extended Data Fig. 2c). Using 589 the five-day threshold and the detrended data, we identified 511 distinct surface MHWs in 590 OISST and 248 bottom MHWs in GLORYS. Many years had multiple MHWs. GLORYS had 591 fewer MHWs partly because the time-series is shorter and partly because the MHWs it recorded 592 were longer in duration (leading to fewer discrete MHW events relative to OISST, which 593 recorded many shorter MHWs).

We then calculated different MHW metrics: MHW cumulative intensity (the anomaly above the 95<sup>th</sup> percentile threshold summed over the duration of the event in °C-days, duration (number of days), and mean intensity (the average anomaly above the 95<sup>th</sup> percentile threshold over the course of the MHW in °C). We chose MHW cumulative intensity for the main analysis because it encompassed elements of both intensity and duration (i.e., cumulative intensity is higher for longer or for more intense MHWs)—but our biomass results did not change if we used
different metrics (Extended Data Fig. 2d, e). Because under our definition all MHWs exceeded
95% of temperature anomalies in the region, even MHWs with relatively low cumulative
intensities represented extreme events. Unless otherwise specified, models in the main text and
Supplement used MHW cumulative intensity centered and scaled within regions, although our
results were not sensitive to whether we scaled and centered within *vs* across regions (Supp. Tab.
5).

In coral reef ecology, a threshold for ecological damage (i.e., coral bleaching) has been 606 607 identified using degree heating days-the number of days that exceed average temperatures for the hottest summer month by at least 1 °C.<sup>16</sup> We also processed the non-detrended GLORYS 608 609 data using this method for each region (Extended Data Fig. 2f). One data product, Coral Reef 610 Watch, calculates the average temperatures for the hottest summer month during a baseline period of 1985-1990 plus 1993.<sup>16</sup> Because GLORYS began in 1993, we used the four-year 611 612 interval 1993-1996 as the baseline to calculate the average temperatures for the hottest summer 613 month.

The five-day threshold for a MHW used in the main analysis was based on empirical analyses demonstrating that contemporary heatwaves last on average 4.6 days<sup>68</sup>, and is widely used in the literature<sup>6</sup>, although we note that alternate methods exist to define and measure MHWs<sup>15,39,69</sup>. We also assessed whether MHW responses would emerge from classifying any daily anomaly (without the five-day cutoff) as a MHW. We found no relationship (Extended Data Fig. 2a).

620 Species and Community Temperature Indices. The Species Temperature Index (STI)
621 and Community Temperature Index (CTI) are measures of thermal affinity at the species and

community level, respectively.<sup>70</sup> We quantified STI as the median sea surface temperatures
found throughout a species' modeled range from the publicly available STI dataset in Burrows et
al.<sup>26</sup>. STI values were available for 844 of our 1772 focal taxa, comprising 82% of total biomass
in our survey dataset. CTI was calculated in each region and year as the biomass-weighted mean
of all STIs, and we used the difference in CTI from one year to the next as our metric of CTI
change. We also quantified the thermal bias of each species relative to the community (STI CTI)<sup>70</sup>.

629 **Community beta-diversity metrics.** Only observations identified at the species level 630 were included in species composition change (beta diversity) analyses. To assess the impact of 631 MHWs on community structure, we compared Bray-Curtis dissimilarity between surveys spanning a MHW to those between years that did not span a MHW. We partitioned Bray-Curtis 632 633 dissimilarity into two components (biomass gradient and balanced variation) using the betapart package in R<sup>30,71</sup>. The biomass gradient component focuses on changes in biomass of species 634 635 between years within the survey region, while the balanced variation component focuses on the 636 substitution of the biomass of one species by the biomass of another species. For comparison, we 637 also calculated occurrence-based dissimilarity metrics (i.e. species presence-absence data) using 638 Jaccard dissimilarity partitioned into nestedness and turnover components. Balanced variation 639 and turnover both measure substitution of species between communities while biomass gradient 640 and nestedness both measure how species are subsetted between communities.

Additional predictors. We conducted supplementary analyses to explore the role of a number of additional predictors of fish biomass change. In addition to the geographical shifts that may lead to changes in biomass and community composition in a fixed area, marine fishes may shift deeper in response to warming<sup>72,73</sup>. We tested for this effect by calculating depth log 645 ratios that described whether assemblages had shifted deeper or shallower from one survey to the 646 next. Depth log ratio was quantified by: 1. Taking an average of depths at which a species was 647 found in each survey and year, using the depth observations for each haul, and weighted by 648 biomass in the haul; 2. Taking a biomass-weighted mean of all species-level depth values for the 649 entire survey; 3. Calculating the log ratio of the survey-level, biomass-weighted depth values 650 from one year to the next. We found no relationship between MHW cumulative intensity and 651 depth log ratio and no difference between depth changes that did and did not follow a MHW 652 (Extended Data Fig. 6, Supp. Tab. 9).

Marine communities across latitudes have responded differently to climate change, with some declines in species richness recorded in the tropics and at equatorward range edges<sup>24,74</sup> and some increases in species richness recorded in colder oceans and at poleward range edges<sup>74,75</sup>. We tested for latitudinal trends in biomass log ratios and found that the direction or magnitude of biomass change was not related to the median latitude of the region (Supp. Tab. 8).

We explored whether species traits helped to predict species-level biomass change in general, and specifically in the context of MHWs. All fish species traits were obtained from the database in Beukhof et al.<sup>51</sup>. Of the 1772 taxa used in the main analysis, 1620 had trophic level data, 1591 had feeding mode data, and 1612 had habitat data. The pattern of no relationship between MHW cumulative intensity and biomass log ratio persisted when data were grouped by trophic level, feeding mode, or habitat (Extended Data Fig. 7).

664 Some studies find that marine communities respond rapidly to environmental change<sup>76</sup>. 665 Others suggest that ecological responses may lag disturbances by years<sup>77</sup>. We explored whether 666 MHW data from further into the past—up to five years before each trawl survey—predicted 667 biomass responses. Analogous to our findings for MHWs that occur up to 12 months before each survey reported in the main text, we found no evidence that biomass change is associatedgenerally with MHW cumulative intensity from prior years (Supp. Tab. 6).

670 Because fishing, through increased mortality, can influence temporal biomass change, we 671 also analyzed the effects of catch on biomass change. We extracted a historical time-series of reconstructed catch values from the Sea Around Us database<sup>78</sup> by Marine Ecoregions (MEs).<sup>79</sup> 672 The Sea Around Us reconstructed catches are spatially allocated to half degree ocean cells,<sup>80</sup> 673 674 which permits catch data to be assigned to spatial entities such as the 232 MEs identified by Spalding et al.<sup>79</sup> We then paired our survey footprints with the most-overlapping ME. In most 675 676 cases, the MEs and survey footprints were similar. For two large surveys (the West Coast and the 677 Northeast US) we summed catch data across two adjacent MEs. Two small surveys (France and 678 the English Channel) did not correspond well to the MEs and were omitted from the fishing 679 analysis. Because catch data are recorded by calendar year, and the surveys often occur midyear, 680 we fitted models comparing biomass change in a given year to the mean catch level in the past three calendar years (Supp. Tab. 10). 681

682 Statistical methods. We tested for the effects of MHWs using linear models, generalized 683 linear models, or generalized additive models for continuous variables. Models and transformation of variables are described in Supp. Tab. 2-11. Generalized linear models were 684 fitted with the R package "glmmTMB"<sup>81</sup> and generalized additive models with the R package 685 "mgcv"82. When comparing MHW versus non-MHW effects we used two-sided t-tests. While 686 687 not all of the datasets were normally distributed, the t-test is insensitive to skewness for large sample sizes such as ours, whereas non-parametric alternatives are better suited to smaller 688 sample sizes<sup>83</sup>. 689

690 **Power analysis.** We simulated data to assess whether our study had sufficient power to 691 detect MHW-driven biomass changes. We fitted an autoregressive linear model of log biomass 692 over time (Gompertz model) to each region's biomass data, including MHW presence/absence as 693 a predictor. We extracted the coefficient  $\rho$ , intercept  $\alpha$ , and conditional standard deviation  $\sigma$  of 694 this model, and used them to simulate data from the same Gompertz model

$$\ln(B_t) = \alpha + \rho \times \ln(B_{t-1}) + \gamma \times MHW_t + \sigma'$$

696 where B represents biomass in year t, MHW is a binary variable for MHW presence/absence, and  $\gamma$  represents the "true" MHW effect that we varied to explore power. This simulation also 697 698 included an error term  $\sigma'$  calculated as a random draw from a normal distribution with mean 0 699 and standard deviation  $\sigma$ . We (1) varied the number of years the simulation was run (assuming 700 that each of the 18 surveys was conducted for that number of years) from 10 to 40 in 1-year steps and 50-200 in 10-year steps with a fixed value of  $\gamma = \ln(0.94)$ , corresponding to the 6% loss of 701 702 biomass predicted by Cheung et al.<sup>3</sup>; and (2) varied  $\gamma$  to represent biomass losses ranging from 703 1% to 30% (in 1% increments up to 10%, and then in 5% increments) given the actual number of 704 years of data we have (n = 369 for GLORYS and n = 441 for OISST). Note that the mean survey 705 duration in our analysis was 20-25 years depending on the paired temperature dataset used. For 706 all these scenarios, simulations were run for each individual survey, converted into log ratio units 707 (as used in the main text), scaled and centered within regions, and pooled across regions. Each 708 set of simulations was run 1000 times for each condition (survey and either number of years or 709 γ).

With these four simulated datasets—a true MHW effect on biomass of -6% and variable
numbers of years, or a fixed number of years from the real dataset and a variable effect of MHW
on biomass, each for OISST and GLORYS—we conducted the same statistical tests as we did in

713	the main text to test for an effect. For every iteration of the simulation, we split the biomass log		
714	ratio data into MHW and non-MHW years and compared the two with a two-sided t-test. We		
715	then calculated what proportion of those tests were significant ( $p = 0.05$ ). These results are		
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717			
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787

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802

## 803 <u>Author contributions</u>

804	All authors contributed to writing and revising the manuscript. ALF, LP, WWLC, MLP, AAM,		
805	ZJK, MLDP, JTT, AA, BM, JPA, and NS contributed to the study conception and design. ALF,		
806	LP, MLP, AAM, ZJK, TLF, MLDP, JTT, BM, and JPA contributed to data acquisition and		
807	analysis. All authors approved the submitted manuscript and subsequent revisions.		
808			
809	Competing interest declaration		
810	The authors declare no competing interests.		
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812	Additional information		
813	Supplementary Information is available for this paper. Correspondence and requests for materials		
814	should be addressed to ALF. Reprints and permissions information is available at		
815	www.nature.com/reprints.		
816			
817	Code availability statement		
818	The code for this study is publicly available on GitHub at		
819	https://github.com/afredston/marine_heatwaves_trawl and archived at		
820	https://doi.org/10.17605/OSF.IO/H6UKT.		
821			
822	Data availability statement		
823	The data used in this project are available on <u>https://doi.org/10.17605/OSF.IO/H6UKT</u> .		
824			
825	Extended data figures		
826			



828 Extended Data Fig. 1. Alternate version of Fig. 2 from the main text, showing results by

region. MHWs were calculated from the detrended GLORYS sea bottom temperature data with
a five-day minimum duration threshold for MHWs, as used in the main text. Points represent log
ratios of mean biomass in a survey from one year to the next. The fitted lines are linear
regressions. The shaded areas are 95% confidence intervals. Survey names and sample sizes per
survey are listed in Supp. Tab. 1.





# 836 Extended Data Fig. 2. Results did not change when alternative methods were used to





- detrended data, (d) using a MHW metric of duration (days), (e) using a MHW metric of intensity
- 840 (°C), (f) calculating degree heating days instead of MHW anomalies, and (g) using only summer

841 MHWs (see Methods). The fitted lines are linear regressions. The shaded areas are 95%







Extended Data Fig. 3. Marine heatwave cumulative intensity (total anomaly in °C-days) in
each survey region with and without detrending the temperature data to remove the signal

847 of secular warming. The main text results are detrended. Here, we plot MHW cumulative

848 intensity based on all SBT anomalies from GLORYS, rather than applying the five-day threshold

that was used the main text, to more clearly show the differences between the two methods.



852 Extended Data Fig. 4. Daily 95<sup>th</sup> percentile anomalies in the two marine heatwave data

853 sources: sea surface temperature from OISST and sea bottom temperature from GLORYS

854 (both detrended). To simplify comparison we plot all anomalies, not just those MHWs that

exceeded a five-day threshold. Note that the OISST time-series began in 1982 and GLORYS

began in 1993. Region names are listed in Supp. Tab. 1.





**community.** We calculated mean abundance (a), mean biomass (b, used in the main text),

861 median abundance (c), and median biomass (d). MHWs were calculated from the detrended

862 GLORYS sea bottom temperature data with a five-day minimum duration threshold for MHWs,

as used in the main text. Points represent log ratios of each metric in a survey from one year to

the next (n = 343). The fitted lines are linear regressions. The shaded areas are 95% confidence

865 intervals. The Northeast US survey was omitted because it did not have abundance data

866 recorded.

867







870 heatwaves. Fish assemblage depth change (log ratio) was not predicted by (a) the presence or

absence of a MHW or (b) MHW cumulative intensity (total anomaly in °C-days; n = 369).

872 MHWs were calculated from the detrended GLORYS sea bottom temperature data with a five-

873 day minimum duration threshold for MHWs, as used in the main text. The fitted line in (b) is a

874 linear regression and the shaded area is its 95% confidence interval.





877 Extended Data Fig. 7. Marine heatwave effect on taxon-specific biomass log ratios grouped

by traits. Biomass log ratio and MHW cumulative intensity (total anomaly in °C-days) grouped by (a) feeding mode (n = 29,628), (b) trophic level (n = 29,909), and (c) habitat preference (n = 29,681) of each taxon. Trait data were extracted from Beukhof et al.<sup>51</sup> (see Methods). MHWs were calculated from the detrended GLORYS sea bottom temperature data with a five-day

- 882 minimum duration threshold for MHWs, as used in the main text. Fitted lines are linear
- regressions. Shaded areas are 95% confidence intervals.
- 884





886 Extended Data Fig. 8. The presence or absence of a MHW did not affect temporal

**community dissimilarity.** We measured community dissimilarity as partitioned occurrence-



and partitioned biomass-based beta diversity metrics of substitution and subset (Bray-Curtis balanced variation (c) and biomass gradient (d)). Community dissimilarity metrics were calculated within each region from one year to the next (n = 369). MHWs were calculated from the detrended GLORYS sea bottom temperature data with a five-day minimum duration threshold for MHWs, as used in the main text.





Extended Data Fig. 9. Results from a power analysis simulating how much data would be
required to detect a range of MHW-induced biomass losses. Approximately 600 survey-years
in total (summed across all regions) would be required to find a significant effect if MHWs

- reduced biomass by 6% using either the GLORYS (a) or OISST (b) datasets; the dashed vertical
- 900 line shows the sample size of our actual datasets. Given the true size of our datasets (n = 369
- 901 survey-years for GLORYS and 441 for OISST), our analysis had the power to detect a MHW-
- 902 induced biomass decline of ~9% with GLORYS (c) and ~8% with OISST (d). The dashed
- 903 horizontal line denotes one conventionally accepted threshold for power (0.8).



906	<b>Extended Data Fig.</b>	10. Biomass trends over time in each su	urvey. The top five taxa by

- 907 biomass are highlighted. Shaded grey rectangles denote when any MHWs occurred in the
- 908 preceding survey-year. MHWs were calculated from the detrended GLORYS sea bottom
- 909 temperature data with a five-day minimum duration threshold for MHWs, as used in the main
- 910 text. Note that x- and y-axes vary depending on time-series length and overall survey catch.